

# **Anomaly Detection in Space-Time (and higher dimensional) Point Processes**

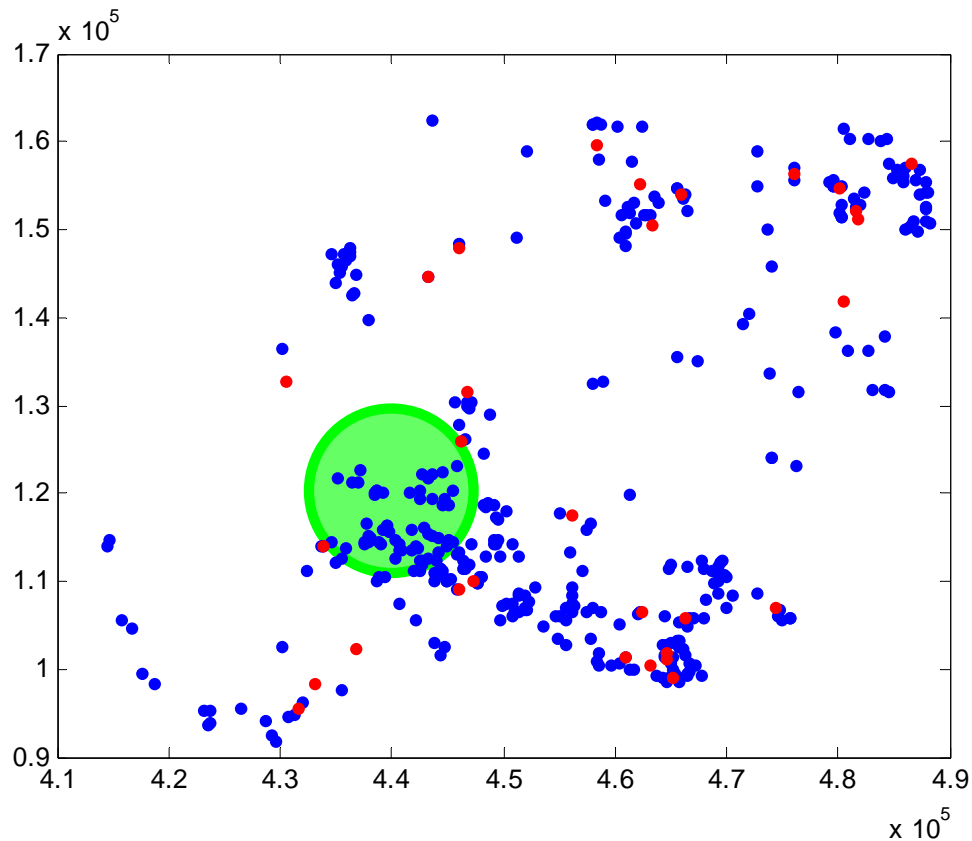
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# AEGISS Data

*Diggle, P, & Rowlingson, B, & Su, T.L. (2005) Environmetrics*

## AEGISS - Ascertainment and Enhancement of Gastrointestinal Infection Surveillance and Statistics



Blue: 1 Jan 2003 – 10 Feb 2003

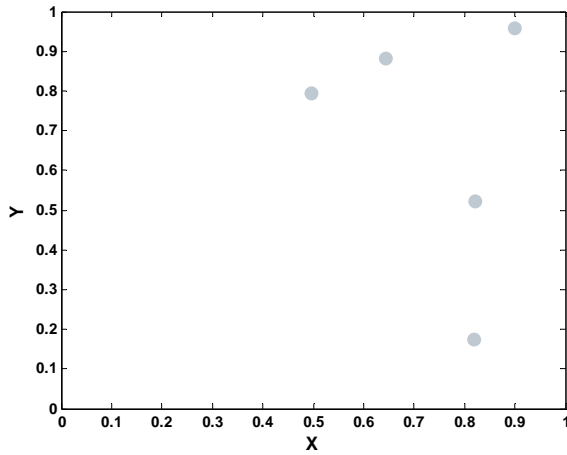
Red: 11 Feb 2003 – 16 Feb 2003

# Problem Description

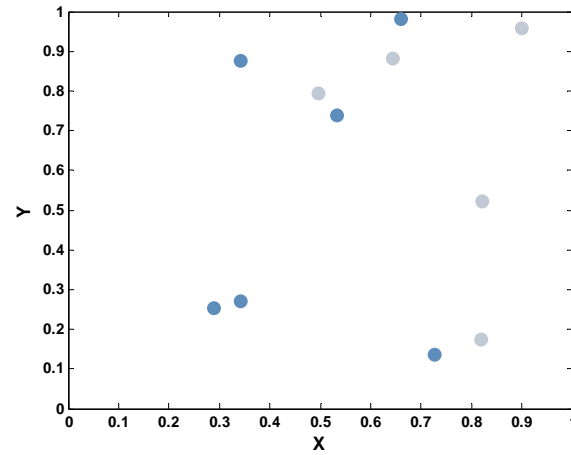
- **Quickly detect (space-time)**
  - Criminal/Terrorist outbreaks or change in tactics
  - Disease Outbreaks (Spatial and Syndromic)
  - Computer/Network Intrusions
  - Targets in Sensor Systems
- **Terminology: Change Detection**
  - Prospective Change Detection
  - Real-Time Change Detection
  - On-Line Change Detection
  - Surveillance
  - Stopping Time
  - Poisson Disorder Problem

# Type of Change Considered

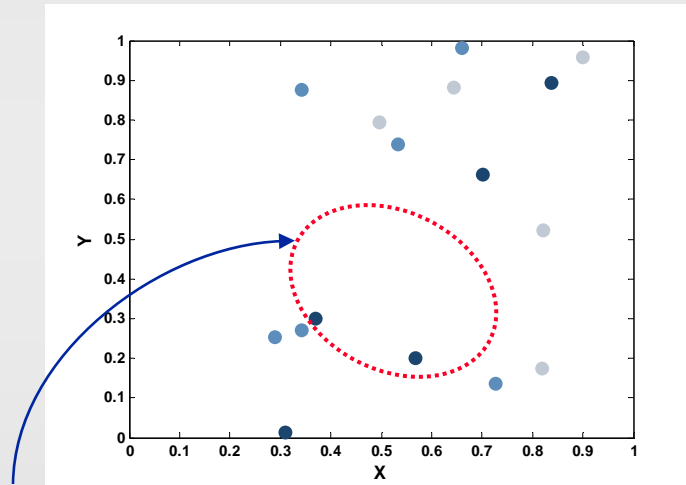
- **Anomaly (a specific type of change)**
  - A localized (in space) region of unusual activity – in this case a change in intensity measure of the point process
- **Anomaly Detection**
  - Detection of when an anomaly occurs and the specific region where it occurs
- **Spatial Anomaly vs. Time series**
- **Methodological Goal**
  - ‘Quick Detection of Anomalies in Space-Time Point Processes’



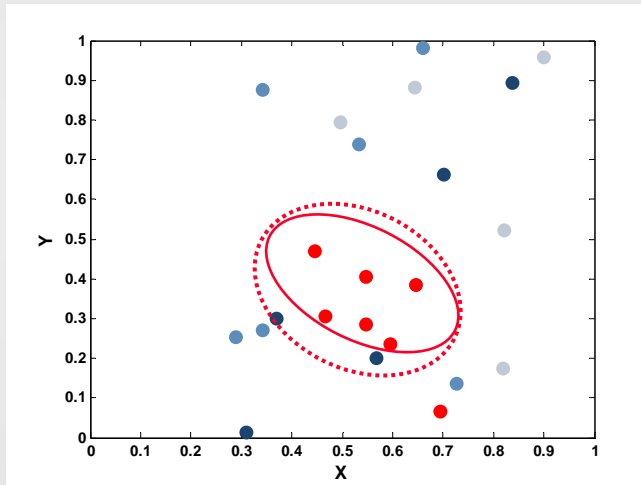
$T=1, N(1)=5$



$T=2, N(2)=6$

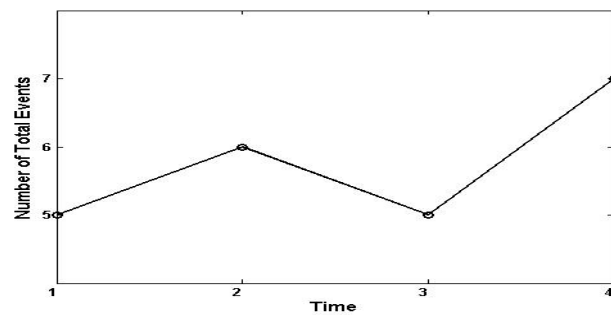


$T=3, N(3)=5$



$T=4, N(4)=7$

Or change  
at time  $T=3$ ,  
not  
detected  
until  $T=4$ ?



# Problem Formulation

- $N_t(\mathbf{B})$  is space-time Poisson process with intensity  $\lambda_0(t, \mathbf{B})$ 
  - $t \in \mathbf{Z}^+$  (Discretize time)
  - $\mathbf{B} \subseteq \mathbf{W} \subset \mathbf{R}^d$  (or more generally  $\mathbf{R}^k \times \mathbf{C}^l$ )

$$H_o : N_t(B) \sim \lambda_0(t, B) \quad \{t \in \mathcal{R}^+, B \subseteq W\}$$

$$H_a : \exists B^*, t^* :$$

$$\text{S.T. } N_t(B^*) \sim \delta \cdot \lambda_0(t, B^*) \quad \{t \geq t^*, \delta \neq 1\}$$

Change parameters:

$t^*$  - Temporal change point

$B^*$  - Spatial change region

$\delta$  – change rate

# GLR for Space-Time Anomalies

- **Lorden's Generalized Likelihood Ratio (GLR)**

- When unknown post-change parameter

$$U_t^* = \max_{1 < k < t} \left[ \sup_{\theta_1 \in \Theta_1} \prod_{i=k}^t \Lambda(i, \theta_1, \theta_0) \right]$$

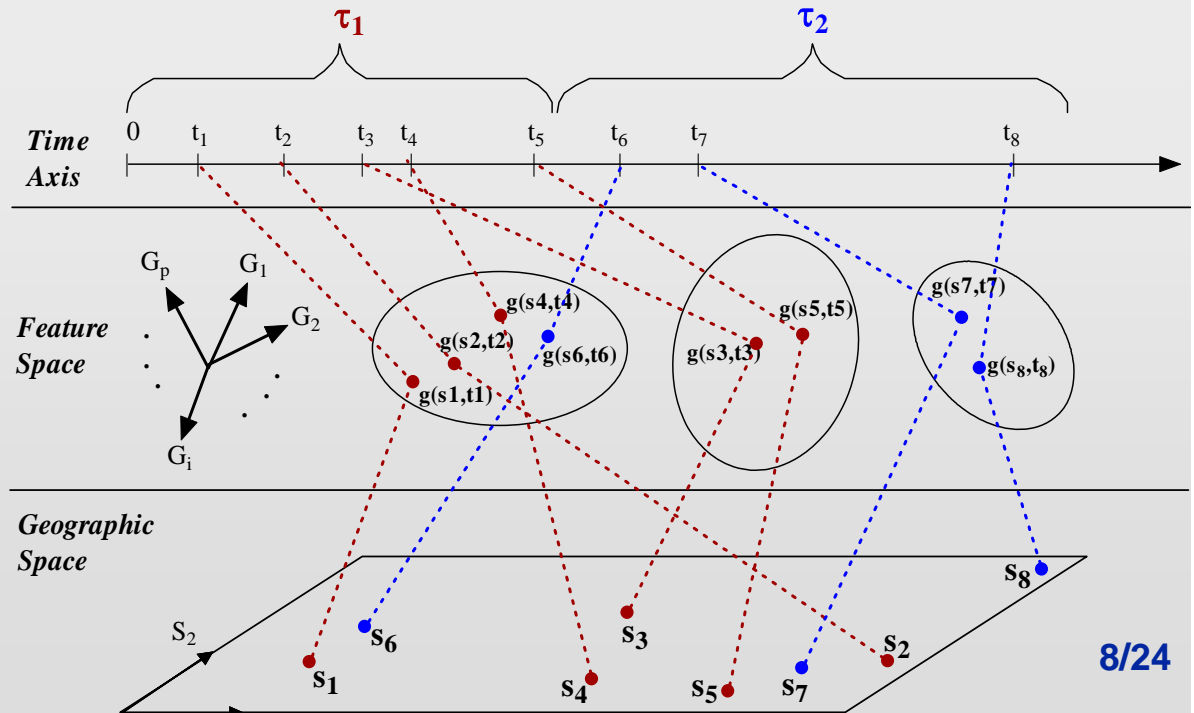
Alternative:  $\theta_1 = [B, \delta, \lambda_0(B)]$

Null:  $\theta_0 = \lambda_0(B)$

- For each possible change point  $k$ , estimate the parameters under the alternative model, and calculate the likelihood ratio.
- Choose the value of  $k$  which maximizes the LR
- Space-Time Scan Statistic (Kulldorff, JRSS-A 2001)

# Finding Most Likely Change Region

- At every time  $t$ , and possible change point  $k$ , scan for the (connected) region  $B$  that maximizes the  $U_t^*$  statistic
- Scan in geographic space or feature space?
- Examples
  - Disease - space or features (covariates, risk factors)
  - Computer Intrusion – space or features
  - Site selection - feature
  - Sensors  $\mathbb{R}^2$  or  $\mathbb{R}^3$



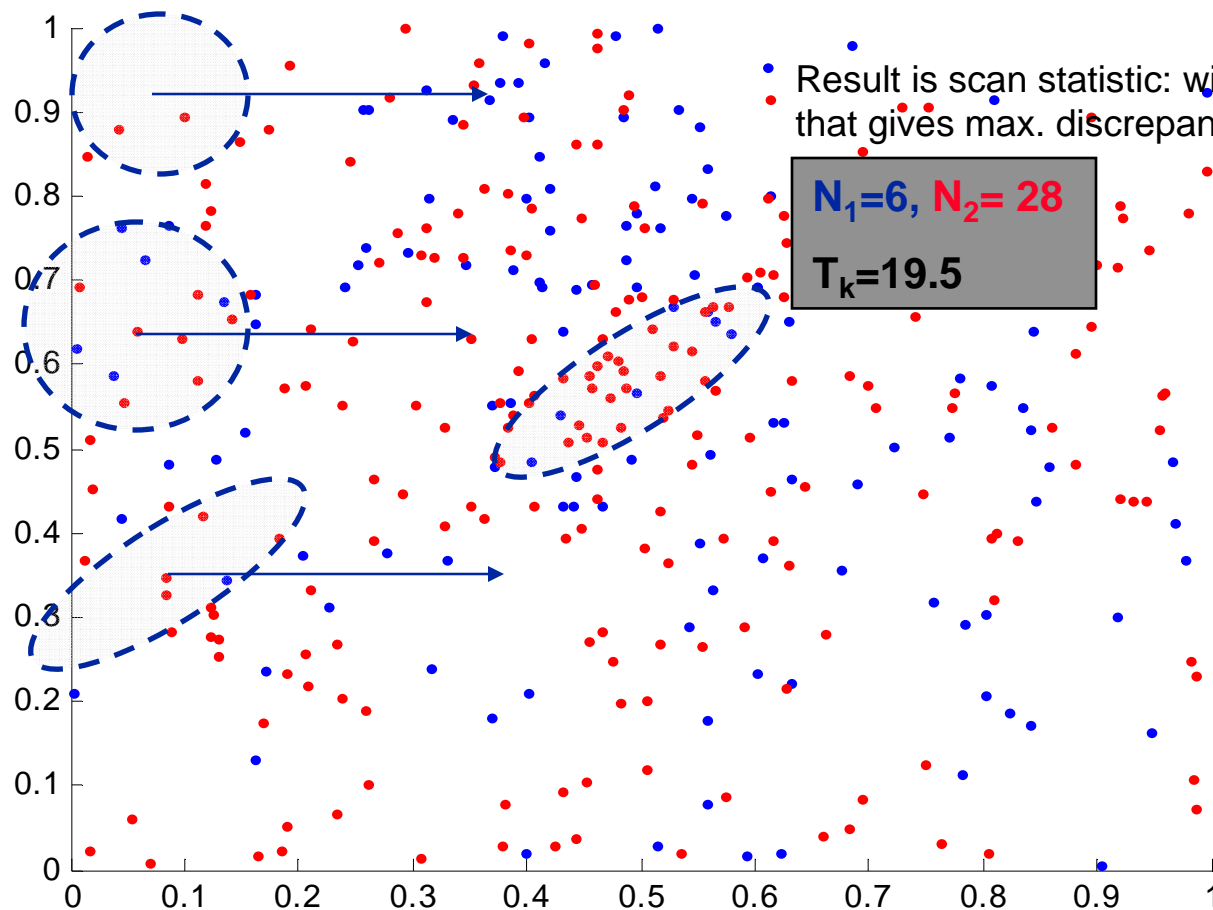


# Generalized Scan Process

- A scan process can be used to detect a (connected) region of unusual observations
- Create a window,  $W_s$  of some geometry and move it over the entire region of interest ( $\forall s \in W$ ) and calculate some discrepancy score for each window,  $D(W_s)$ .
- Testing will be carried out on the *scan statistic*,  $T = \max_{s \in W} D(W_s)$ .
  - Can include multiple window geometries
- The GLR statistics becomes  $U_t^* = \max_{1 < k < t} T_k$
- Problems:
  - Dependent upon the geometries and sizes of the windows
  - Gets difficult in high dimensions

# Generalized Scan Process

Choose windows and scan over entire region (for each window)



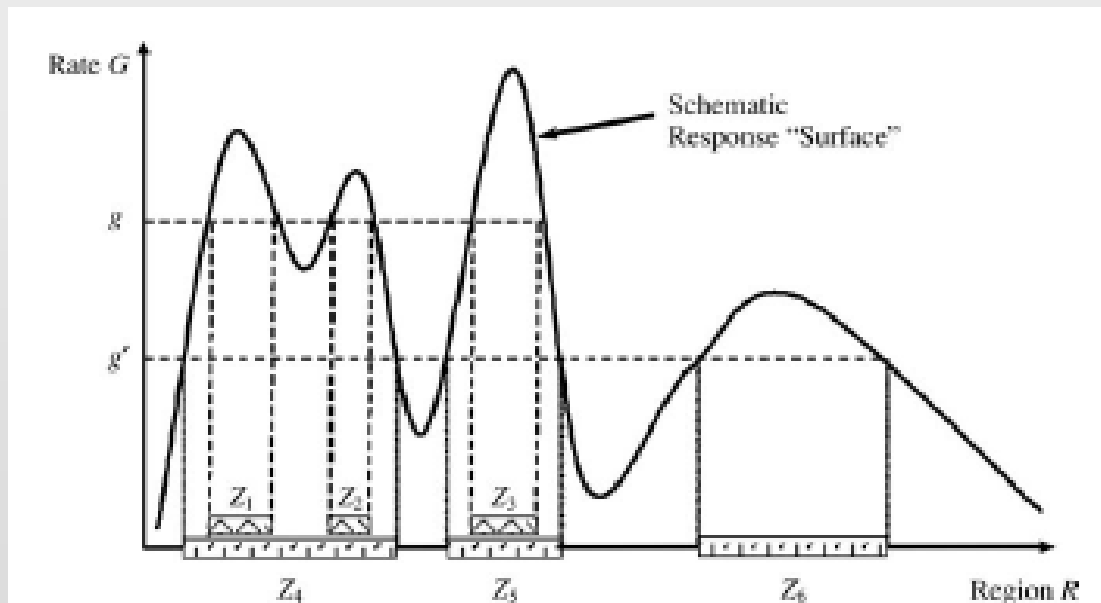
Result is scan statistic: window and location that gives max. discrepancy score

$N_1=6$ ,  $N_2=28$

$T_k=19.5$

# Alternatives to Scan Process

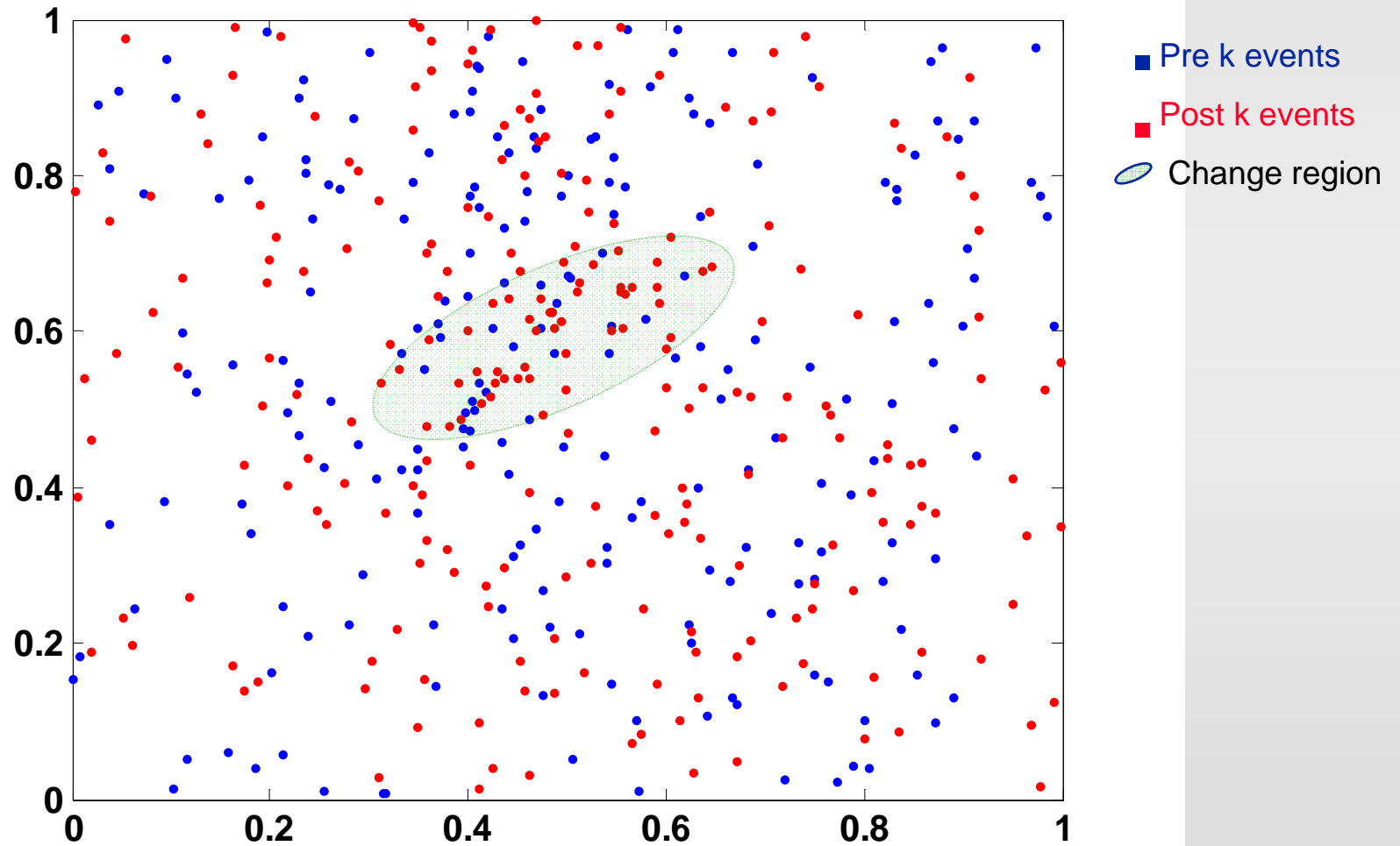
- **Random Searches** (Porter and Brown, CSDA 2007)
  - Modified PRIM for high dimensions (mixed variables)
  - Maximum of Random Partition Search
- **Upper Level Set (ULS) Approaches** (Patil and Tallie, Env. & Ecol. Stats 2004)
  - Kernel intensity estimation
  - Random Partition Search



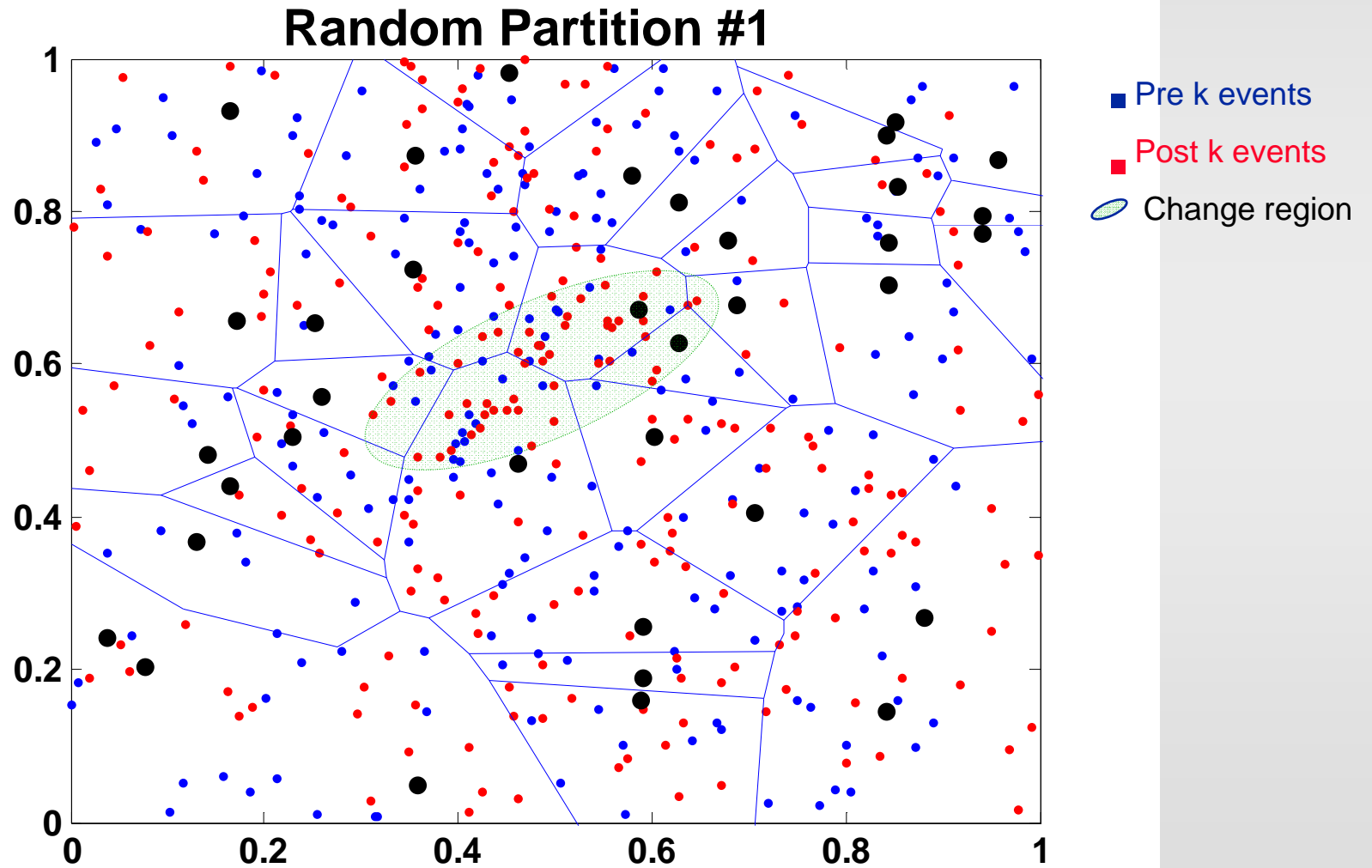
# Random Voronoi Partition Search

- **Mix of random searching and ULS to find region of maximum discrepancy**
- **Generate M Poisson Voronoi partitions of W**
  - Generators are Poisson distributed with rate  $\mu$
  - For each partition, select the cell with the max discrepancy
- **Identify  $B^*$  as either:**
  - The max cell over all M
  - Score the region in each max cell with a value of 1, search through ULS to maximize  $U_t^*$ 
    - » Here, only considering connected regions.

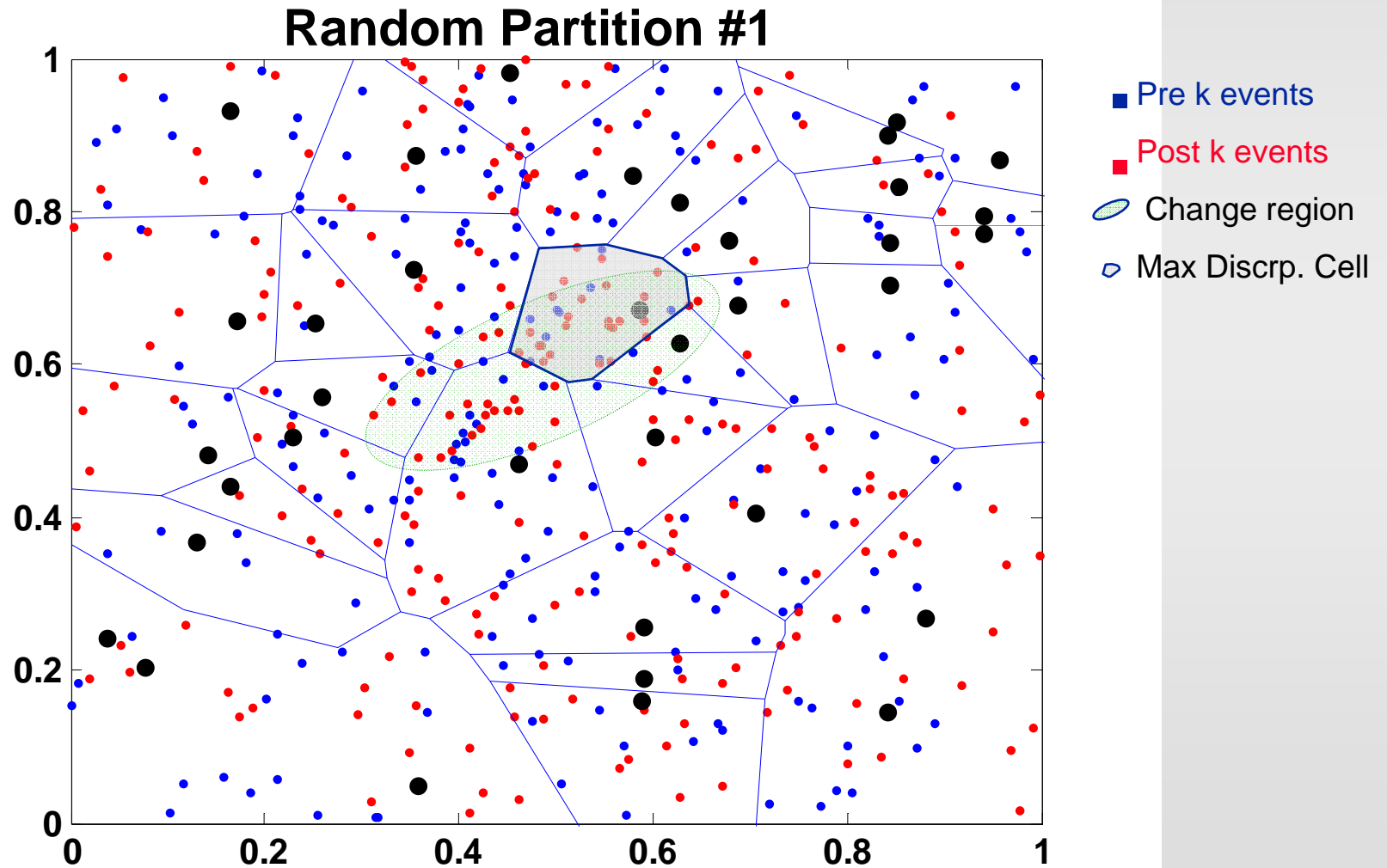
# Realization of Space-Time PP



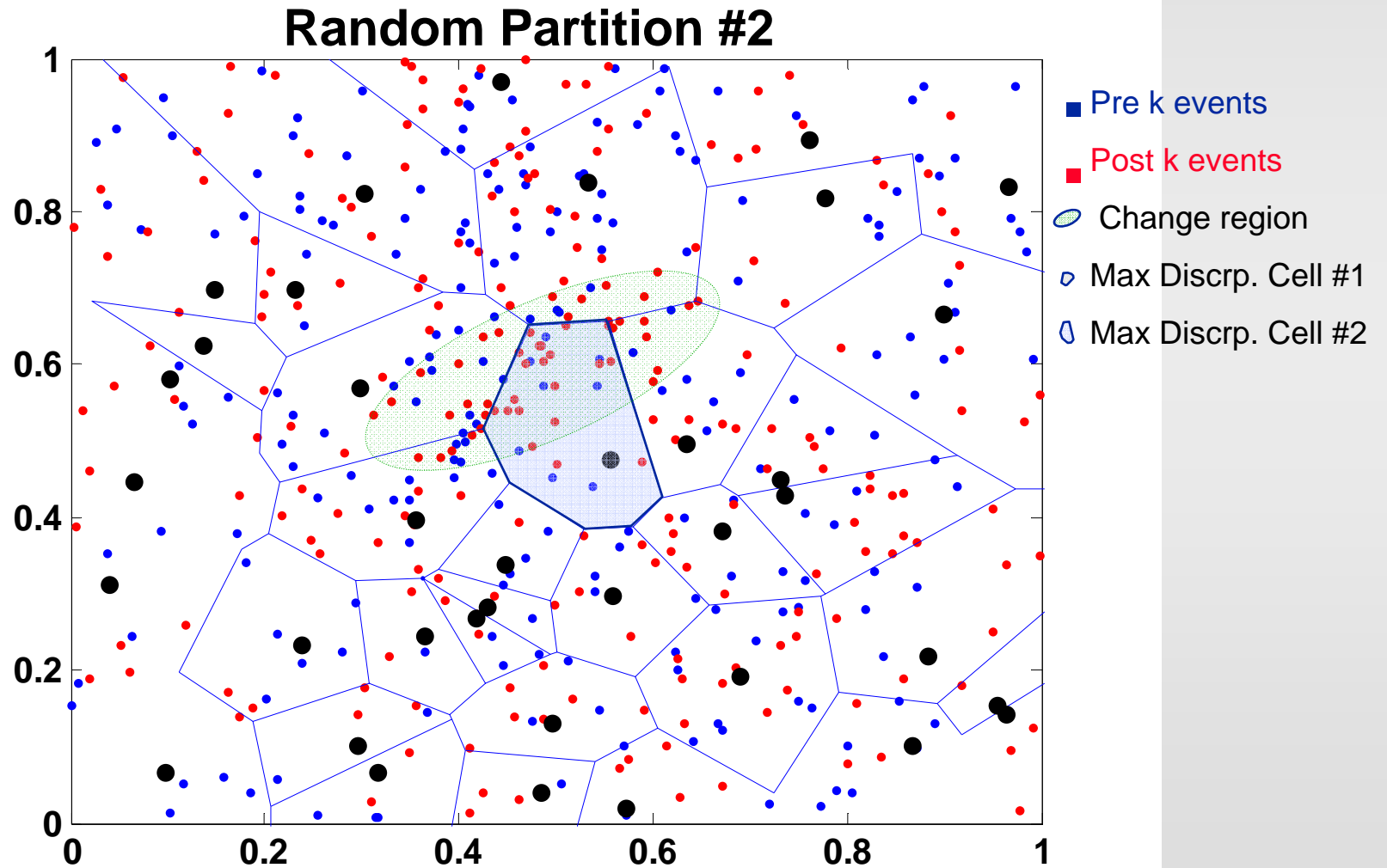
# Random Voronoi Partition Search



# Random Voronoi Partition Search

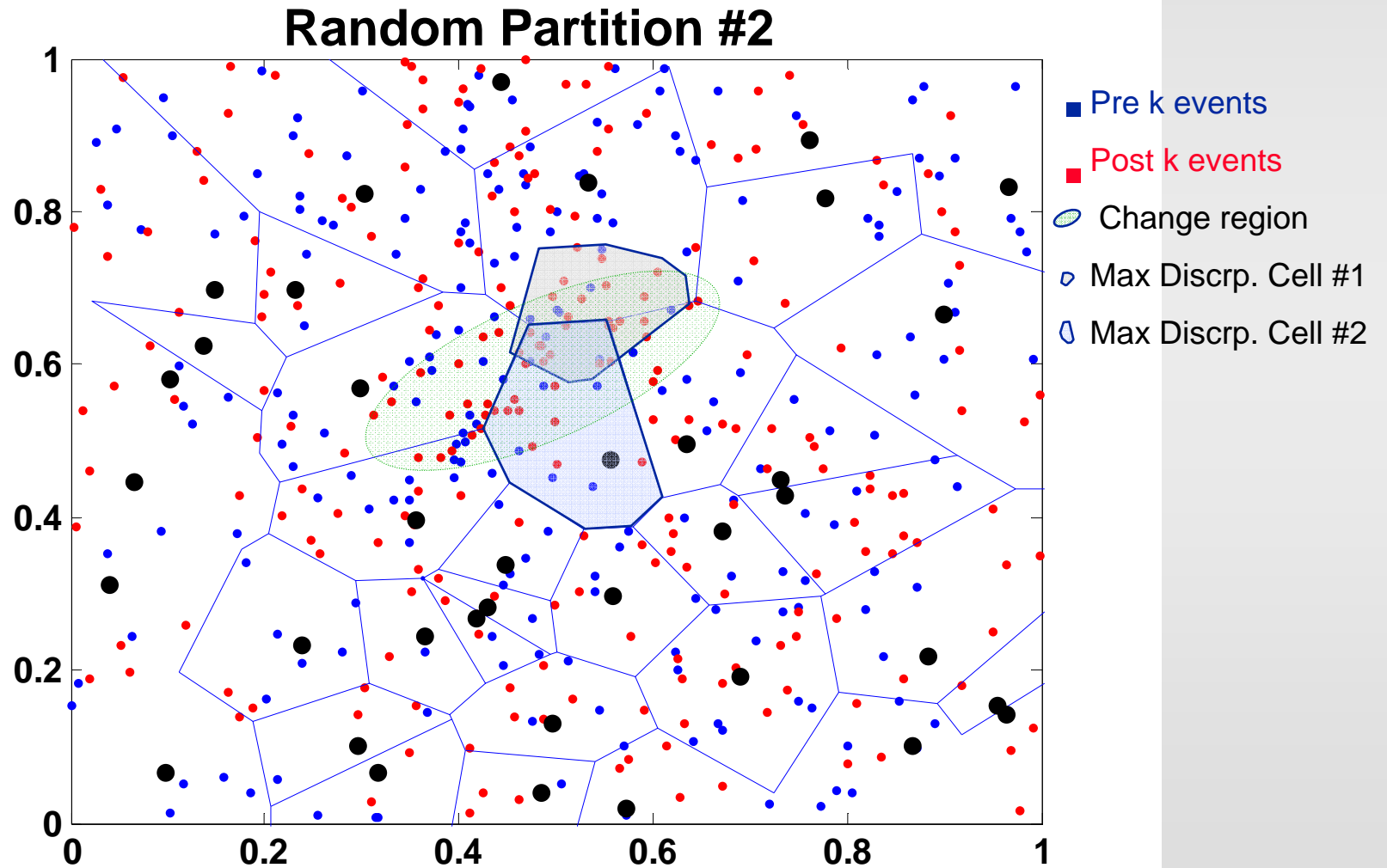


# Random Voronoi Partition Search

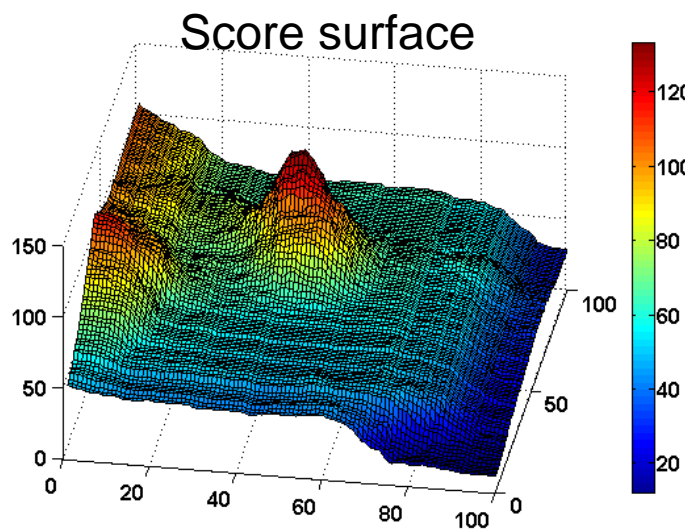
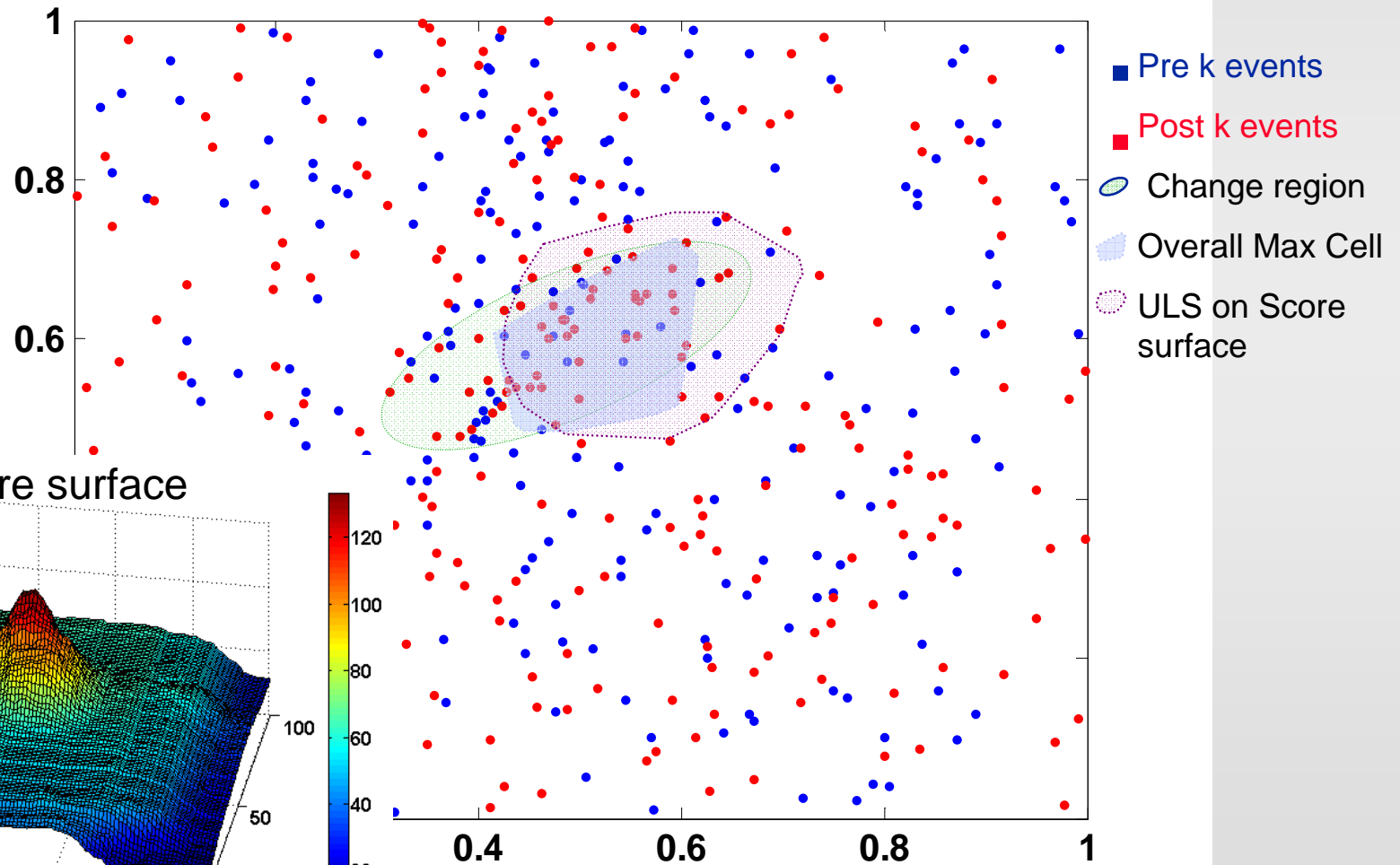




# Random Voronoi Partition Search



# Random Voronoi Partition Search



# Problems with GLR

- **GLR can become very large, even under the null model**
  - Especially if allowed to search over large class of shapes
- **Anomalous region difficult to estimate precisely**
- **This can result in an increased false alarm rate as  $U_t^*$  will take large values even if no change**
- **Difficult to obtain performance bounds**
  - Distribution of  $U_t^*$  unknown
- **Mitigate these problems through *adaptive* procedures**

# Adaptive Procedures

Dragalin, EQC 1997; Lorden & Pollak, Annals 2005; Tartakovski, et al, Stat. Meth. 2006

- **Uses parameter estimates from previous data only**
  - At time  $t-1$ , an estimate of  $\theta_1, \theta_0$  is made
  - These values are used in the likelihood ratio at time  $t$

- **Adaptive CUSUM (AdC)**

$$\tilde{U}_t^* = \max_{1 < k \leq t} \prod_{i=k}^t \Lambda(i, \theta_1^{i-1}, \theta_0^{i-1})$$

- **Adaptive SR (ASR)**

$$\tilde{R}_t^* = \sum_{k=1}^t \prod_{i=k}^t \Lambda(i, \theta_1^{i-1}, \theta_0^{i-1})$$

where  $\theta^{t-1}$  is an estimate of  $\theta$  based only on observations  $\{N_1, \dots, N_{t-1}\}$


# Adaptive Properties

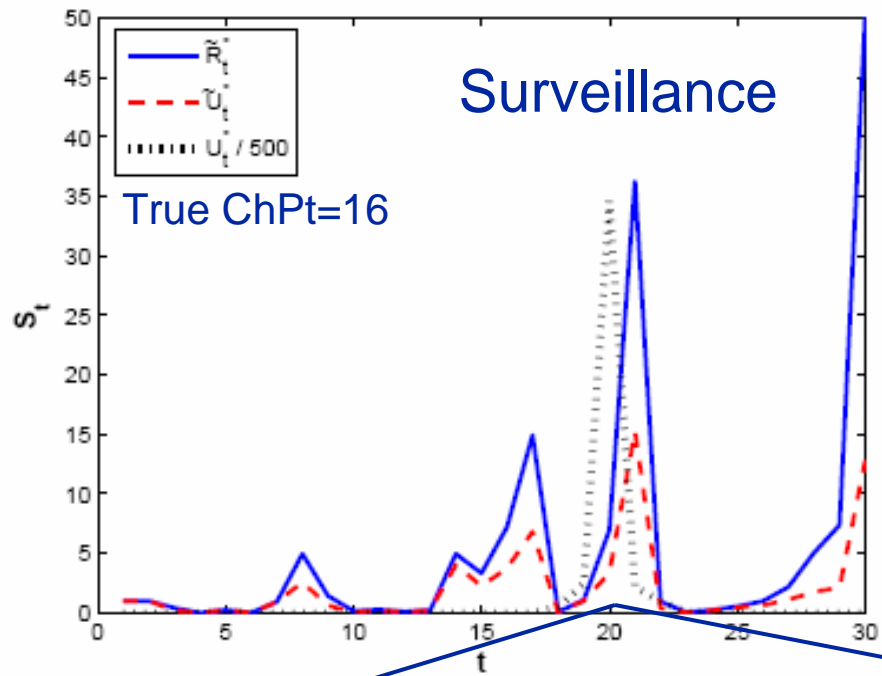
- $E[\Lambda(t, \theta_1^{t-1}, \theta_0) | \mathcal{H}_0] = 1$   
for any value of  $\theta_1^{t-1}$ .
- If  $t < t^*$  then multiplying the  $\Lambda(t, \theta_1^{t-1}, \theta_0^{t-1})$  should remain relatively small
- However, if  $t \geq t^*$  then  $\theta_1^{t-1} = (\mathbf{B}^{t-1}, \delta^{t-1}, \mathbf{k}^{t-1})$  are the ML estimates of the post change parameters
  - $\tilde{U}_t^*$  and  $\tilde{R}_t^*$  will accumulate quickly

# Estimating Adaptive Parameters

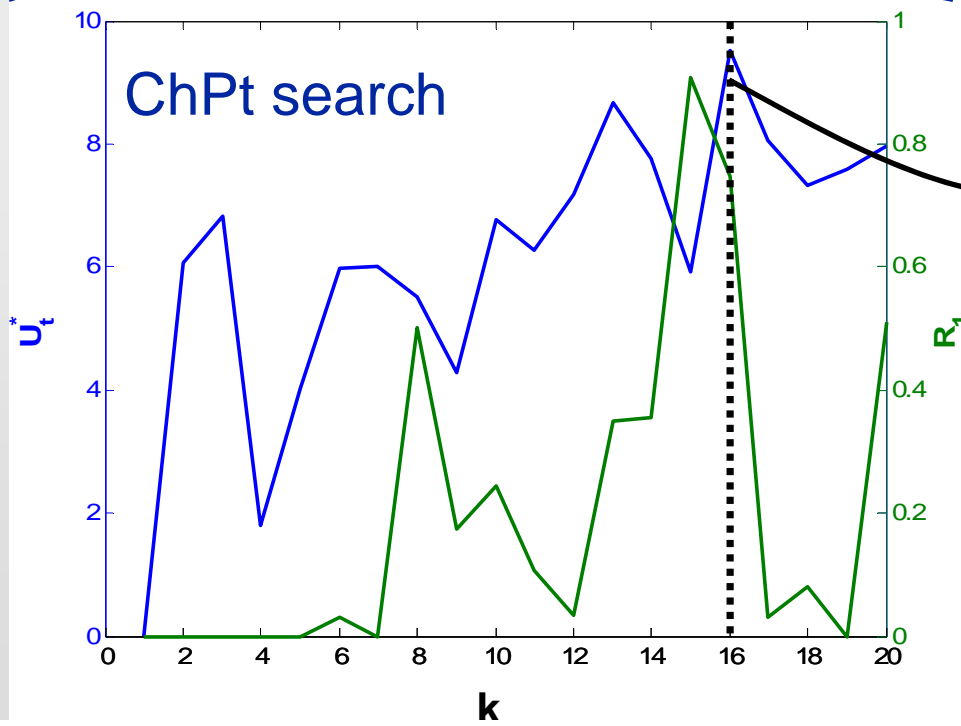
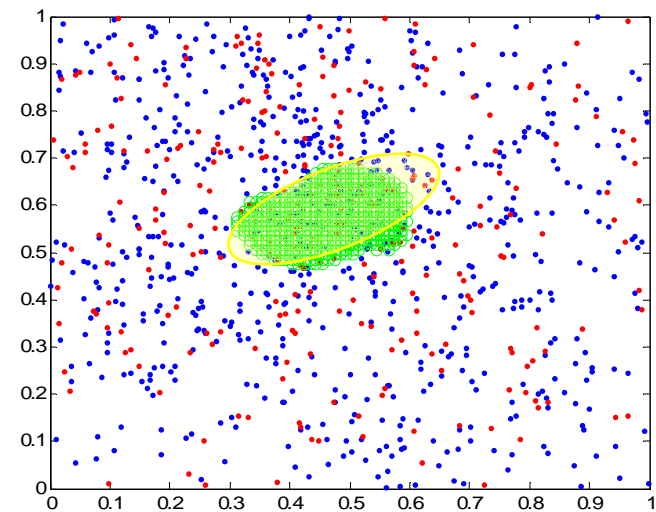
- At time  $t-1$ , estimate  $(\hat{t}^*, \hat{B}, \hat{\delta}, \hat{\lambda}_0(\hat{B}))$  and use these as the parameter estimates at time  $t$

$$(\hat{t}^*, \hat{B}^*, \hat{\delta}) = \arg \left( \max_{0 < k \leq t-1} \left[ \sup_{B \in \mathcal{B}_W} \sup_{\delta \geq 0} \Lambda_0^{t-1}(k, B, \delta, \lambda_0(B)) \right] \right)$$

Let  $(K, B^{t-1}, \delta^{t-1}) = (\hat{t}^*, \hat{B}^*, \hat{\delta})$ .  Previous data only



### Spatial Change Region



# Other Considerations

- **Setting Alarm decision rules**
  - Threshold controls false alarm rate and detection delay
  - Adaptive procedures have good bounds
- **Performance Measures**
  - Bayes Risk, Neyman Pearson, Minimax
  - $E[\text{detection delay}]$ ,  $E[\text{False Alarm Rate}]$
  - How much of  $B^*$  must be captured?
- **Also for marked PP's (space-time case-control, competing animal habitat, etc)**
- **Comparison to space-time scan statistic (SatScan)**



**-¿QUESTIONS?-**

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